# Math 254A Lecture 13 Notes

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# 1 Observing Macroscopic Quantities From Microscopic States

#### 1.1 Recap

We have a phase space  $(M, \lambda)$  which is a  $\sigma$  finite but not finite measure space. The energy of one particle is  $\varphi : M \to [0, \infty)$ , where min  $\varphi = \operatorname{ess\,min} \varphi = 0$ . Then we know that

$$\lambda^{\times n} \left( \left\{ (p_1, \dots, p_n) \in M^n : \frac{1}{n} \Phi_n(p_1, \dots, p_n) := \frac{1}{n} \sum_{i=1}^n \varphi(p_i) \in I \right\} \right)$$
$$= \exp\left( n \cdot \sup_{x \in I} s(x) + o(n) \right),$$

where

$$s(x) = \inf_{\beta > 0} \{ s^*(\beta) + \beta x \}.$$

We also have the Fenchel-Legendre transform

$$s^*(\beta) = \log \int e^{-\beta\varphi}.$$

 $\beta$  achieves equality in the definition of s

$$\iff s \text{ has a tangent of slope } \beta \text{ at } x$$
$$\iff D_+ s(x) \le \beta \le D_- s(x)$$
$$\iff s^*(\beta + (-s(x)) = -\beta x$$
$$\iff D_- s^*(\beta) \le -x \le D_+ s^*(\beta)$$
$$\iff s^* \text{ has a tangent of slope } -x \text{ at } \beta.$$

Using  $s^*$ , we can prove:

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$$s^*(\beta) \to \begin{cases} \log \lambda(\{\varphi = 0\}) & \beta \to \infty\\ \infty & \beta \downarrow 0. \end{cases}$$

- $s^*$  is strictly decreasing and strictly convex.
- $s^*$  is differentiable on  $(0, \infty)$ .

$$s(x) \to \begin{cases} \log \lambda(\{\varphi = 0\}) & x \downarrow 0\\ \infty & x \to \infty. \end{cases}$$

- *s* is strictly increasing and strictly concave.
- s is differentiable on  $(0, \infty)$ .

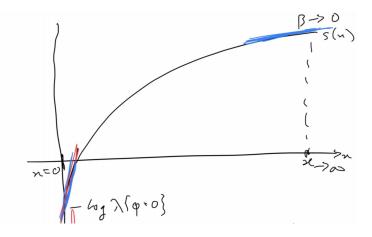
### **1.2** Behavior of s'

Let's analyze the behavior of s':

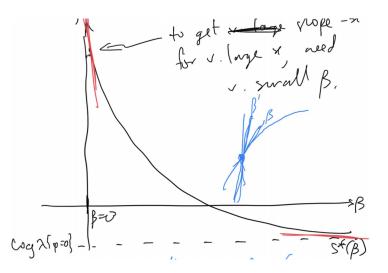
#### Proposition 1.1.

$$s'(x) \to \begin{cases} 0 & x \to \infty \\ \infty & x \to 0. \end{cases}$$

Instead of a formal proof, here are some pictures. Look at the possible slopes we can get for points on the graph of s and how they correspond to slopes for points on the graph for  $s^*$ .



To get slope -x for very large x in the graph of  $s^*$ , we need very small  $\beta$ .



#### 1.3 Observing macroscopic quantities from microscopic states

Now imagine we are looking at some other macroscopic observable quantity of the microscopic state  $(p_1, \ldots, p_n) \in M^n$ . We will study functions for the form

$$\Psi_n(p_1,\ldots,p_n) = \sum_{i=1}^n \psi(p_i).$$

If  $M = \mathbb{R}^3 \times \mathbb{R}^3$ , we could take  $\psi(r, p) = \mathbb{1}_D(r)$ , which indicates whether a particle is in D or not in D; then  $\Psi_n$  would be the total number of particles in D.

We need some regularity. A simple sufficient condition is that  $\psi$  is bounded. A weaker but still sufficient condition is that for every  $\beta > 0$ , there is an  $\varepsilon > 0$  such that  $\int e^{-\beta\varphi} e^{-\gamma\psi} d\lambda < \infty$  for all  $\gamma \in (-\varepsilon, \varepsilon)$ .

Let's assume  $\psi$  is bounded, and we'll ask about the distribution of  $\Psi_n$  on the approximate level set  $\{\frac{1}{n}\Phi_n \in I\}$ , where I is a small interval. We need to compare  $\lambda^{\times n}(\{\frac{1}{n}\Phi_n \in I\})$  and  $\lambda^{\times n}(\{\frac{1}{n}\Phi_n \in I, \frac{1}{n}\Psi_n \in J\})$ . We use the generalized type-counting machinery with  $\mathbb{R}^2$  to get an asymptotic for this:

$$\begin{split} \lambda^{\times n} \bigg( \bigg\{ \frac{1}{n} \Phi_n \in I, \frac{1}{n} \Psi_n \in J \bigg\} \bigg) &= \lambda^{\times n} \left( \bigg\{ (p_1, \dots, p_n) \in M^n : \frac{1}{n} \sum_{i=1}^n (\varphi(p_i), \psi(p_i)) \in I \times J \bigg\} \bigg) \\ &= \exp\bigg( n \cdot \sup_{(x,y) \in I \times J} \widetilde{s}(x, y) + o(n) \bigg), \end{split}$$

where  $\widetilde{s}(x,y): \mathbb{R}^2 \to [-\infty,\infty)$  is an upper semicontinuous, concave function with

$$\widetilde{s}(x,y) = \inf_{\beta,\gamma} \{ \widetilde{s}^*(\beta,\gamma) + \beta x + \gamma y \}$$

and Fenchel-Legendre transform

$$\widetilde{s}^*(\beta,\gamma) = \log \int e^{-\beta\varphi} e^{-\gamma\psi} d\lambda$$

Here, we assume  $\psi$  is bounded,  $|\psi| \leq M$ , so

$$\widetilde{s}^*(\beta,\gamma) = \begin{cases} \infty & \beta = 0 \\ < \infty & \beta > 0. \end{cases}$$

Here,  $\tilde{s}(x, y) \leq s(x)$  for all  $y \in \mathbb{R}$ . We want to find a  $y_0$  such that  $\tilde{s}(x, y_0) = s(x)$  and  $\tilde{s}(x, y) < s(x)$  for any other y. This will tell us that conditioned on  $\Phi$  being x, we are likely to have  $\Psi$  be  $y_0$  and not likely to have any other y. We have

$$s(x) = \inf_{\beta>0} \left\{ \log \int e^{-\beta\varphi} d\lambda + \beta x \right\},$$

which is greater than or equal to

$$\widetilde{s}(x,y) = \inf_{\beta > 0, \gamma \in \mathbb{R}} \left\{ \log \int e^{-\beta \varphi} e^{-\gamma \psi} \, d\lambda + \beta x + \gamma y \right\}$$

**Lemma 1.1.**  $\tilde{s}(x, y_0) = s(x)$  and  $\tilde{s}(x, y) < s(x)$  for any other y, where

$$y_0 = \frac{\int \psi e^{-\beta\varphi} \, d\lambda}{\int e^{-\beta\varphi} \, d\lambda} = \langle \psi, \mu_\beta \rangle$$

and

$$d\mu_{\beta}(p) = rac{e^{-eta arphi(p) \, d\lambda(p)}}{\int e^{-eta arphi} \, d\lambda}$$

is the **Gibbs measure** obtained from  $\lambda, \varphi, \beta$ .

*Proof.* First, s is differentiable, so for every x > 0, there is a unique  $\beta > 0$  such that  $s(x) = \log \int e^{-\beta\varphi} d\lambda + \beta x$ . To achieve  $\tilde{s}(x, y_0) = s(x)$ , we must have that the function  $\gamma \mapsto \log \int e^{-\beta\varphi} e^{-\gamma\psi} d\lambda + \beta x + \gamma y_0$  achieves its minimum uniquely at  $\gamma = 0$ . This function of  $\gamma$  is convex (by Hölder), strictly convex if  $\psi$  is not a.s. constant, and differentiable. Assuming  $\psi$  is not a.s. constant, we need  $y_0$  such that

$$\frac{\partial}{\partial\gamma} \left\{ \log \int e^{-\beta\varphi} e^{-\gamma\psi} \, d\lambda + \beta x + \gamma y_0 \right\} = 0$$

at  $\gamma = 0$ . This is the derivative of the log of the moment generating function. Differentiate under the integral to get

$$\frac{\partial}{\partial\gamma}\log\int e^{-\beta\varphi}e^{-\gamma\psi}\,d\lambda = \left.\frac{\int -\psi e^{-\beta\varphi}e^{-\gamma\psi}\,d\lambda}{\int e^{-\beta\varphi}e^{\gamma\psi}\,d\lambda}\right|_{\gamma=0} = -\langle\psi,\mu_\beta\rangle$$

So  $\frac{\partial}{\partial \gamma} [\cdots]|_{\gamma=0} = -\langle \psi, \mu_{\beta} \rangle + y_0$ , and this equals 0 iff  $y_0 = \langle \psi, \mu_{\beta} \rangle$ .

Corollary 1.1.

$$\lambda^{\times n} \left( \left\{ \left| \frac{1}{n} \Psi_n - \langle \psi, \mu_\beta \rangle \right| > \varepsilon \right\} \ \left| \ \left\{ \frac{1}{n} \Phi_n \in I \right\} \right) \le e^{-c \cdot n + o(n)},$$

where c is a constant, I is a short enough interval containing x, and we are using conditional probability notation.

**Remark 1.1.** Given  $\frac{1}{n}\Phi_n \approx x$ , we found that

$$\Psi_n \approx n (\text{its average over } \{\frac{1}{n} \Phi_n \approx n\}^1)$$
  
$$\approx n \langle \psi, \mu_\beta \rangle$$
  
$$= \langle \psi(p_1) + \dots + \psi(p_n), \mu_\beta^{\times n} \rangle$$
  
$$= \int \Psi_n \, d\mu_{\beta,n},$$

where

$$d\mu_{\beta,n}(p_1,\ldots,p_n) = \frac{e^{-\beta\Phi_n(p_1,\ldots,p_n)} d\lambda^{\times n}(p)}{\int e^{-\beta\Phi_n} d\lambda^{\times n}} = \mu_\beta \times \cdots \times \mu_\beta$$

is called the **canonical ensemble measure**.

<sup>&</sup>lt;sup>1</sup>This is called the **microcanonical ensemble**.